

Math 461 Fall 2021

Renming Song

University of Illinois at Urbana-Champaign

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Outline

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- 1 **General Info**
- 2 6.2 Independent random variable
- 3 6.3 Sums of independent random variables.

HW7 is due Friday, 10/22, before the end of class time . Please submit your HW7 via the course Moodle page.

Solution to HW6 is on my homepage now.

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Two random variables X and Y are said to be independent if for any two subsets A and B of \mathbb{R} ,

$$P(X \in A, Y \in B) = P(X \in A)P(Y \in B).$$

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Example 2

Suppose that the number of people entering a certain post office on a given day is a Poisson random variable with parameter $\lambda > 0$. Assume that each person entering the post office is male with probability p and female with probability $1 - p$, independent of all others. Show that the number of males and the number of females entering the post office on a given day are independent Poisson random variables with parameters λp and $\lambda(1 - p)$ respectively.

Let X and Y be the number of males and the number of females entering the post office on a given day respectively. X and Y are non-negative integer-valued random variables.

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For any non-negative integers i and j ,

$$\begin{aligned}P(X = i, Y = j) &= P(X = i, Y = j | X + Y = i + j)P(X + Y = i + j) \\&= \binom{i+j}{i} p^i (1-p)^j e^{-\lambda} \frac{\lambda^{i+j}}{(i+j)!} \\&= e^{-\lambda p} \frac{(\lambda p)^i}{i!} e^{-\lambda(1-p)} \frac{(\lambda(1-p))^j}{j!}.\end{aligned}$$

Hence

$$P(X = i) = e^{-\lambda p} \frac{(\lambda p)^i}{i!} \sum_{j=0}^{\infty} e^{-\lambda(1-p)} \frac{(\lambda(1-p))^j}{j!} = e^{-\lambda p} \frac{(\lambda p)^i}{i!}.$$

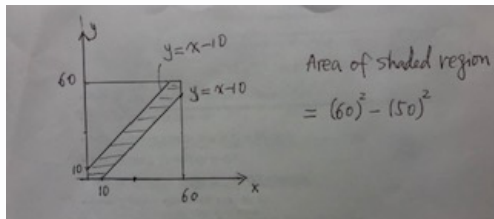
Similarly

$$P(Y = j) = e^{-\lambda(1-p)} \frac{(\lambda(1-p))^j}{j!}.$$

Therefore X and Y are independent Poisson random variables with parameters λp and $\lambda(1-p)$ respectively.

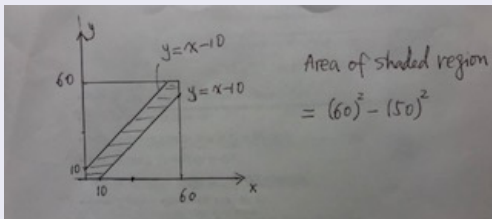
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The answer is

$$\frac{60^2 - 50^2}{60^2} = \frac{11}{36}.$$

Proposition

- (i) Suppose that X and Y are discrete with joint mass function $p(\cdot, \cdot)$. Then X and Y are independent if and only if

$$p(x, y) = g(x)h(y), \quad (x, y) \in \mathbb{R}^2$$

for some functions g and h on \mathbb{R} .

- (ii) Suppose that X and Y are jointly abs. cont. with joint density $f(\cdot, \cdot)$. Then X and Y are independent if and only if

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Example 4

The joint density of X and Y is

$$f(x, y) = \begin{cases} 10e^{-(2x+5y)}, & x > 0, y > 0 \\ 0, & \text{otherwise.} \end{cases}$$

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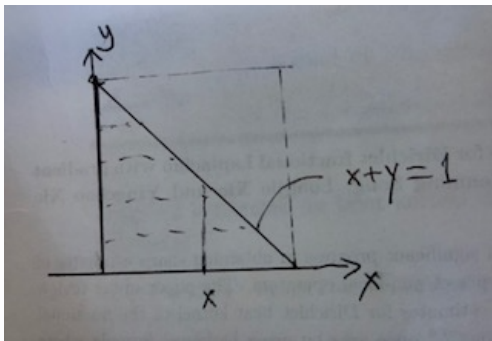
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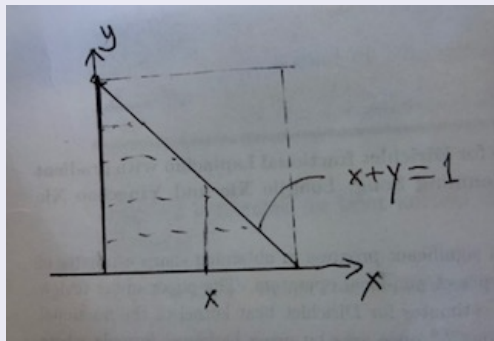
$$f(x, y) = \begin{cases} 24xy, & x \in (0, 1), y \in (0, 1), x + y \in (0, 1) \\ 0, & \text{otherwise.} \end{cases}$$



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Both X and Y take values in $(0, 1)$. For $x \in (0, 1)$,

$$f_X(x) = \int_0^{1-x} 24xydy = 12x(1-x)^2.$$

Similarly, for $y \in (0, 1)$,

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n random variables X_1, \dots, X_n are said to be independent if for any subsets A_1, \dots, A_n of \mathbb{R} ,

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Suppose that X_1, \dots, X_n are independent absolutely continuous random variables with a common density f . Define

$$U = \min\{X_1, \dots, X_n\}, \quad V = \max\{X_1, \dots, X_n\}.$$

Find the densities of U and V respectively.

Let's deal with V first. Let F be the common distribution. For any $v \in \mathbb{R}$,

$$P(V \leq v) = P(X_1 \leq v, \dots, X_n \leq v) = (F(v))^n.$$

Thus the density of V is $f_V(v) = n(F(v))^{n-1}f(v)$.

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Now let's deal with U . For any $u \in \mathbb{R}$,

$$\begin{aligned} P(U \leq u) &= 1 - P(U > u) = 1 - P(X_1 > u, \dots, X_n > u) \\ &= 1 - (1 - F(u))^n. \end{aligned}$$

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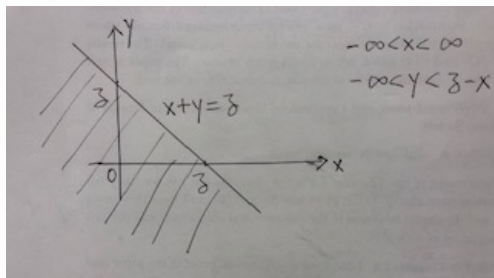
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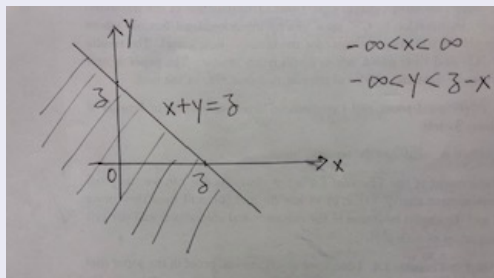
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For any $z \in \mathbb{R}$,

$$\begin{aligned}F_Z(z) &= P(X + Y \leq z) \\&= \int_{-\infty}^{\infty} \left(\int_{-\infty}^{z-x} f_X(x) f_Y(y) dy \right) dx \\&= \int_{-\infty}^{\infty} \left(\int_{-\infty}^z f_X(x) f_Y(v-x) dv \right) dx, \quad (y = v-x) \\&= \int_{-\infty}^z \int_{-\infty}^{\infty} f_X(x) f_Y(v-x) dx dv\end{aligned}$$

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